Video 1. Watch the introduction to hypothesis tests.
We start with a null hypothesis $H_{0}$, which we'll assume to be true until we have evidence to the contrary. Exactly what constitutes contrary evidence is determined by our choice of alternative hypothesis $H_{1}$. The null hypothesis should specify the distribution of a test statistic. We collect data, calculate the observed value of the test statistic, then calculate the $\mathbf{P}$-value (or observed significance level) of our data. We reject $H_{0}$ in favor of $H_{1}$ if the p-value is small (usually less than 0.05 ). Otherwise we fail to reject $H_{0}$.
Method. For tests about a proportion, the null hypothesis should be $H_{0}: \theta=\theta_{0}$ and our test statistic is the sample total $T$. Under $H_{0}, T \sim \operatorname{binom}\left(n, \theta_{0}\right)$.

1. A July, 2018 NPR/IPSOS poll asked respondents if they support or oppose "building a wall or fence along the entire U.S./Mexico border." Let $\theta$ be the proportion of the population in question (e.g. Midwesterners) that oppose building a wall or fence. Test $H_{0}: \theta=0.5$ against $H_{1}: \theta>0.5$ for the following populations. State your p-values and conclusions clearly.
a) 115 of 217 people in the Midwest oppose the wall

Solution. P-value: $P(T \geq 115)=1-P(T \leq 114)=0.2076776$. Fail to reject $H_{0}$.
b) 150 of 264 people in the West oppose the wall

Solution. P-value: $P(T \geq 150)=1-P(T \leq 149)=0.01551882$ Reject $H_{0}$.
c) 198 of 401 people in the South oppose the wall

Solution. P-value: $P(T \geq 198)=1-P(T \leq 197)=0.617747$ Fail to reject $H_{0}$.
2. The article "Analysis of Reserve and Regular Bottlings: Why Pay for a Difference Only the Critics Claim to Notice?" reported on an experiment to determine if wine tasters could correctly distinguish between reserve and regular versions of a wine. In each trial tasters were given 4 indistinguishable containers of wine, two of which contained the regular version and two of which contained the reserve version of the wine. The taster then selected 3 of the containers, tasted them, and was asked to identify which one of the 3 was different from the other 2. In 855 trials, 346 resulted in correct distinctions. Does this provide compelling evidence that wine tasters can distinguish between regular and reserve wines?
a) Start with a null hypothesis that the tasters can't distinguish between the wines. State this as a hypothesis about the proportion of times the tasters correctly identify the odd wine out.

Solution. Let $\theta$ be the proportion of wine tasters who can distinguish between regular and reserve wines. If tasters can't distinguish between the wines, then they'll be correct purely by chance. This will happen with probability $1 / 3$. Thus our null hypothesis is $H_{0}: \theta=1 / 3$.
b) State an alternative hypothesis.

Solution. I think the most reasonable alternative is $H_{1}: \theta>1 / 3$.
c) Calculate the observed significance level of the experimental results.

Solution. $T \sim \operatorname{binom}(855,1 / 3)$. P-value: $P(T \geq 346)=1-P(T \leq 345)=7.600273 \times 10^{-6}$. Reject $H_{0}$. This is very strong evidence that tasters can tell the difference, at least some of the time. They're correct far more often than we'd expect them to be if they're just guessing.
d) What is your conclusion? Any additional comments/thoughts?

Solution. We have rejected our null hypothesis and concluded that wine tasters can distinguish between regular and reserve wines, at least some of the time. Looking at the data shows that we shouldn't conclude that the tasters are good at making the distinction, however. They're correct about $40 \%$ of the time, which is only a bit more often than if they guessed at random ( $33 \%$ of the time). One way to think of this: average tasters drinking 100 bottles of reserve wine would only notice they weren't drinking regular wine 7 times (be careful, though, since this is probably not very reflective any individual's actual experience).

Method. For tests about the mean of a normally distributed population (with unknown variance), the null hypothesis should be $H_{0}: \mu=\mu_{0}$ and our test statistic is $T=\frac{\bar{X}-\mu_{0}}{S / \sqrt{n}}$. Under $H_{0}, T$ has a $t$-distribution with $n-1$ degrees of freedom.
3. The mean unrestrained compressive strength for a random sample of 45 specimens of a particular type of brick was 3107 psi, and the sample standard deviation was 188 psi. Does the data indicate that the true average unrestrained compressive strength is less than the design value of 3200 ? Assume the population is normally distributed and test $H_{0}: \mu=3200$ against $H_{1}: \mu<3200$ at a significance level of 0.05 .
Solution. Test statistic: $t=\frac{3107-3200}{188 / \sqrt{45}}=-3.31842$. This comes from a $t$-distribution with 44 degrees of freedom. The p-value is $P(T \leq-3.31842)=0.0009119218$. Therefore we should reject $H_{0}$ and conclude that the bricks are not as strong as they're designed to be.

Definition. When conducting a hypothesis test there are two types of error:
(1) Type I error is rejecting $H_{0}$ when $H_{0}$ is actually true.
(2) Type II error is failing to reject $H_{0}$ when $H_{0}$ is actually false.

The significance level of a test is the probability of a type I error.
Ideally, the probability of both types of error would be small. Unfortunately, decreasing the probability of type I error (by lowering the threshold for significance from 0.05 to 0.01 , say) increases the probability of type II error (and vice versa).
4. The EPA has determined that the Maximum Contaminant Level Goal (MCLG, "The level of a contaminant in drinking water below which there is no known or expected risk to health") for nitrates in drinking water is $10 \mathrm{mg} / \mathrm{L}$. Imagine you have been hired by the City of Spokane to monitor drinking water safety. Your plan is to collect a random sample of water from 25 different sources and test the samples for nitrate levels, then use the sample mean as an estimate of mean level in all of Spokane's water.
a) State your null and alternative hypotheses in both non-mathematical language and as $H_{0}: \mu=$ $\qquad$ with a corresponding $H_{1}$.
Solution. We have two options for the hypotheses:
(a) Assume the water is safe until we find evidence it isn't: $H_{0}: \mu=10$ with $H_{1}: \mu>10$;
(b) Assume the water isn't safe until we find evidence it is: $H_{0}: \mu=10$ with $H_{1}: \mu<10$.
b) State in plain language what a Type I error would be for your hypotheses (e.g. if you were explaining your results to the Mayor).

Solution. The two setups above have different type I errors:
(a) Conclude the water isn't safe when it actually is;
(b) Conclude the water is safe when it actually isn't.
c) Repeat for a Type II error.

Solution. The two setups above have different type II errors:
(a) Continuing to think the water is safe when it actually isn't;
(b) Continuing to think the water isn't safe when it actually is.
d) What significance level do you think makes sense for the test?

Solution. The significance level is the probability of a type I error. For the first setup, I want to make the probability of type II error as small as possible, since this error leads to people drinking unsafe water. This means making the significance level as large as is as acceptable, maybe 0.1.

For the second setup, type I error leads to people drinking unsafe water, so I want this to be small. Maybe a significance level of 0.001 would be acceptable.
e) The EPA statement on the effect of nitrate contamination: "Infants below the age of six months who drink water containing nitrate in excess of the MCL could become seriously ill and, if untreated, may die. Symptoms include shortness of breath and blue-baby syndrome." Does this change the significance level you want to use for your test? Does this change how you want to set up the null and alternative hypotheses?
Solution. This suggests that the second setup is better, since I really want to minimize the chance of children drinking unsafe water. I might also want to make the significance level even smaller.

