

PROBABILITY

Postulate 1. $P(A) \geq 0$

Postulate 2. $P(S) = 1$

Postulate 3. If A_1, A_2, \dots are mutually exclusive events, then $P(A_1 \cup A_2 \cup \dots) = P(A_1) + P(A_2) + \dots$

Definition. A and B are **independent** if and only if $P(A \cap B) = P(A)P(B)$.

Definition. If $P(B) > 0$, then the conditional probability of A given B is $P(A|B) = \frac{P(A \cap B)}{P(B)}$

Proposition. If $P(A > 0)$ and $P(B > 0)$, then $P(B|A) = \frac{P(A|B)P(B)}{P(A)}$

Proposition. If B_1, B_2, \dots, B_k partition S , then $P(A) = \sum_{i=1}^k P(A|B_i)P(B_i)$.

Definition. The cdf of a random variable X is $F(x) = P(X \leq x)$.

Definition. Random variables X and Y are *independent* if and only if $f(x, y) = f_X(x)f_Y(y)$ for all possible x, y .

IMPORTANT SUMS

Proposition. $(x + y)^n = \sum_{r=0}^n \binom{n}{r} x^{n-r} y^r$

Proposition. $\sum_{n=0}^{\infty} u^n = \frac{1}{1-u}$ if $|u| < 1$

Proposition. $\sum_{n=0}^{\infty} \frac{u^n}{n!} = e^u$

Proposition. $\sum_{i=1}^n i = 1 + 2 + \dots + n = \frac{1}{2}n(n+1)$

Proposition. $\sum_{i=1}^n i^2 = 1^2 + 2^2 + \dots + n^2 = \frac{1}{6}n(n+1)(2n+1)$

CENTRAL LIMIT THEOREM

Theorem. If X_1, X_2, \dots, X_n constitute a random sample from an infinite population with mean μ and variance σ^2 , then the limiting distribution of $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$ as $n \rightarrow \infty$ is the standard normal distribution.

Proposition. If \bar{X} is the mean of a random sample of size n from a normal population with mean μ and variance σ^2 , then \bar{X} has a normal distribution with mean μ and variance $\frac{\sigma^2}{n}$.

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Proposition. If \bar{X} is the mean of a random sample of size n from a finite population of size N with mean μ and variance σ^2 , then \bar{X} has mean μ and variance $\frac{\sigma^2}{n} \left(\frac{N-n}{N-1} \right)$.

Definition. The r^{th} **moment about the origin** of a random variable X is $\mu'_r = E(X^r)$.

Definition. The first moment about the origin of a random variable is called the **mean** and is denoted by μ .

Proposition. If a and b are constants, then $E(aX + b) = aE(X) + b$.

Definition. The r^{th} **moment about the mean** of a random variable X is $\mu_r = E[(X - \mu)^r]$.

Definition. The second moment about the mean of a random variable is called the **variance** and is denoted by σ^2 . The **standard deviation** of a random variable is $\sigma = \sqrt{\sigma^2}$.

Proposition (A calculating formula for the variance). $\sigma^2 = \mu'_2 - \mu^2 = E(X^2) - [E(X)]^2$

Proposition. If a and b are constants, then $\text{var}(aX + b) = a^2 \text{var}(X)$.

Proposition (A calculating formula for μ_3). $\mu_3 = \mu'_3 - 3\mu'_2\mu + 2\mu^3$

Definition. The **moment-generating function** of a random variable X is $M_X(t) = E(e^{tX})$.

Proposition. If $M_X(t)$ is the moment-generating function of a random variable X , then $M_X^{(r)}(0) = \mu'_r = E(X^r)$

Proposition. If a and b are constants, then $M_{aX+b}(t) = e^{bt} M_X(at)$.

Definition. If X and Y are jointly distributed random variables with means μ_X and μ_Y , respectively, then $E[(X - \mu_X)(Y - \mu_Y)]$ is called the **covariance** of X and Y and is denoted σ_{XY} , $\text{cov}(X, Y)$, or $C(X, Y)$. If σ_X and σ_Y are the standard deviations of X and Y , respectively, then $\text{cor}(X, Y) = \frac{\text{cov}(X, Y)}{\sigma_X \sigma_Y}$ is the **correlation** of X and Y .

Proposition (A calculating formula for the covariance). $\text{cov}(X, Y) = E(XY) - E(X)E(Y)$

Proposition. If X and Y are independent random variables, then $E(XY) = E(X)E(Y)$.

Proposition. If X_1, X_2, \dots, X_n are random variables and a_1, a_2, \dots, a_n are constants, then

$$E\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i E(X_i)$$

and

$$\text{var}\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i^2 \text{var}(X_i) + 2 \sum_{i < j} a_i a_j \text{cov}(X_i, X_j)$$

Proposition. If X_1, X_2, \dots, X_n are independent random variables and $Y = \sum_{i=1}^n X_i$, then $M_Y(t) = \prod_{i=1}^n M_{X_i}(t)$.

Definition. A random variable X has a **discrete uniform distribution** if it is equally likely to assume any one of a finite set of possible values.

Definition. A random variable X has a **Bernoulli distribution** with parameter θ (with $0 < \theta < 1$) if its probability distribution is

$$f(x; \theta) = \begin{cases} 1 - \theta & \text{if } x = 0 \\ \theta & \text{if } x = 1 \end{cases}$$

The outcome 1 is often referred to as “success” while 0 is “failure” and the experiment is often called a Bernoulli trial.

Proposition. The mean and variance of a Bernoulli random variable are $\mu = \theta$ and $\sigma^2 = \theta(1 - \theta)$.

Definition. The total number of successes in n independent, identically distributed (iid) Bernoulli trials is a random variable with a **Binomial distribution**. A random variable X has a binomial distribution with parameters n and θ if its probability distribution function is

$$b(x; n, \theta) = \binom{n}{x} \theta^x (1 - \theta)^{n-x} \text{ for } x = 0, 1, \dots, n$$

Proposition. The mean and variance of a binomial distribution are $\mu = n\theta$ and $\sigma^2 = n\theta(1 - \theta)$. The moment-generating function of a binomial distribution is $M_X(t) = [1 + \theta(e^t - 1)]^n$.

Definition. Let X_1, X_2, \dots be a sequence of independent, identically distributed (iid) Bernoulli trials, all with probability of success θ . Let N be the trial on which the first success occurs. The random variable N is said to have a **geometric distribution** with parameter θ and its probability distribution function is

$$g(n; \theta) = \theta(1 - \theta)^{n-1} \text{ for } n = 1, 2, 3, \dots$$

Proposition. The mean and variance of a geometric distribution are $\mu = \frac{1}{\theta}$ and $\sigma^2 = \frac{1}{\theta} \left(\frac{1}{\theta} - 1 \right)$.

Definition. Let X_1, X_2, \dots be a sequence of independent, identically distributed (iid) Bernoulli trials, all with probability of success θ . Let N be the trial on which the k^{th} success occurs (so the possible values for N are $k, k+1, k+2, \dots$). The random variable N is said to have a **negative binomial (or binomial waiting-time or Pascal) distribution** with parameters k and θ and its probability distribution function is

$$b^*(n; k, \theta) = \binom{n-1}{k-1} \theta^k (1 - \theta)^{n-k} \text{ for } n = k, k+1, k+2, \dots$$

Proposition. The mean and variance of a negative binomial distribution are $\mu = \frac{k}{\theta}$ and $\sigma^2 = \frac{k}{\theta} \left(\frac{1}{\theta} - 1 \right)$.

Definition. A random variable with the probability distribution function

$$p(x; \lambda) = \frac{\lambda^x e^{-\lambda}}{x!} \text{ for } x = 0, 1, 2, \dots$$

is said to have a **Poisson distribution** with parameter $\lambda > 0$.

Proposition. The mean and variance of a Poisson distribution are $\mu = \lambda$ and $\sigma^2 = \lambda$. The moment-generating function of a Poisson random distribution is $M_X(t) = e^{\lambda(e^t - 1)}$.

Definition. Suppose n elements are to be selected without replacement from a population of size N of which M are successes. The number of successes selected is a **hypergeometric** random variable and its probability distribution function is

$$h(x; n, N, M) = \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}$$

Proposition. The mean and variance of a hypergeometric distribution are $\mu = \frac{nM}{N}$ and $\sigma^2 = \frac{nM(N-M)(N-n)}{N^2(N-1)}$.

Definition. A random variable X has a **uniform continuous distribution** with parameters α and β (with $\alpha < \beta$) if and only if the following function is a probability density for X :
$$f(x; \alpha, \beta) = \begin{cases} \frac{1}{\beta - \alpha} & \text{if } \alpha < x < \beta \\ 0 & \text{elsewhere} \end{cases}.$$

Definition. The **gamma function** is defined as
$$\Gamma(\alpha) = \int_0^\infty y^{\alpha-1} e^{-y} dy \quad \text{for } \alpha > 0.$$

Proposition. For any $\alpha > 1$, $\Gamma(\alpha) = (\alpha - 1)\Gamma(\alpha - 1)$.

Proposition. For any positive integer n ,
$$\Gamma(n) = (n - 1)!$$

Definition. A random variable X has a **gamma distribution** with parameters $\alpha > 0$ and $\beta > 0$ if and only if the following function is a probability density for X :
$$g(x; \alpha, \beta) = \begin{cases} \frac{1}{\beta^\alpha \Gamma(\alpha)} x^{\alpha-1} e^{-\frac{x}{\beta}} & \text{for } x > 0 \\ 0 & \text{elsewhere} \end{cases}.$$

Proposition. A gamma distribution with parameters α and β has moment-generating function
$$M_X(t) = (1 - \beta t)^{-\alpha},$$
 mean
$$\mu = \alpha\beta,$$
 and variance
$$\sigma^2 = \alpha\beta^2.$$

Definition. A random variable X has an **exponential distribution** with parameter $\theta > 0$ if and only if the following function is a probability density for X :
$$g(x; \theta) = \begin{cases} \frac{1}{\theta} e^{-\frac{x}{\theta}} & \text{if } x > 0 \\ 0 & \text{elsewhere} \end{cases} \quad (\text{a gamma distribution with } \alpha = 1 \text{ and } \beta = \theta).$$

Proposition. An exponential distribution with parameter θ has mean
$$\mu = \theta$$
 and variance
$$\sigma^2 = \theta^2.$$

Definition. A random variable X has a **chi-square distribution** with parameter $\nu > 0$ if and only if the following function is a probability density for X :
$$f(x) = \begin{cases} \frac{1}{2^{\nu/2} \Gamma(\nu/2)} x^{\frac{\nu-2}{2}} e^{-\frac{x}{2}} & \text{if } x > 0 \\ 0 & \text{elsewhere} \end{cases} \quad (\text{a gamma distribution with } \alpha = \frac{\nu}{2} \text{ and } \beta = 2).$$

Proposition. A chi-square distribution with parameter ν has mean
$$\mu = \nu$$
 and variance
$$\sigma^2 = 2\nu.$$

Definition. A random variable X has a **normal distribution** with parameters μ and $\sigma > 0$ if and only if the following function is a probability density for X :
$$n(x; \mu, \sigma) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} \quad \text{for all } x \in \mathbb{R}.$$

Proposition. A normal distribution with parameters μ and σ has a moment-generating function
$$M_X(t) = e^{\mu t + \frac{1}{2}\sigma^2 t^2},$$
 mean
$$\mu = \mu$$
 and variance
$$\sigma^2 = \sigma^2.$$

Definition. A random variable X has a **Pareto distribution** with parameters $\alpha > 0$ and $\beta > 0$ if and only if the following function is a probability density for X :
$$f(x) = \begin{cases} \alpha\beta^\alpha x^{-(\alpha+1)} & \text{if } x > \beta \\ 0 & \text{elsewhere} \end{cases}$$